## Essays in Econometrics

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University of Chicago Booth School of Business

Dissertation Defense <sup>1</sup> April 19, 2021

<sup>&</sup>lt;sup>1</sup>Committee: Chris Hansen (chair), Max Farrell, Panos Toulis, Constantine Yannelis

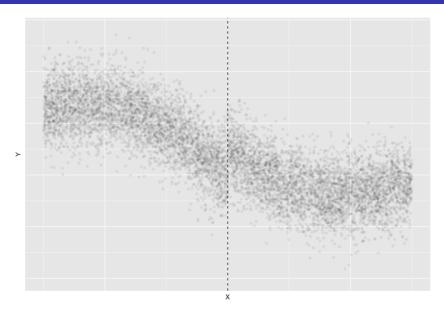
#### Outline

#### 3 Essays

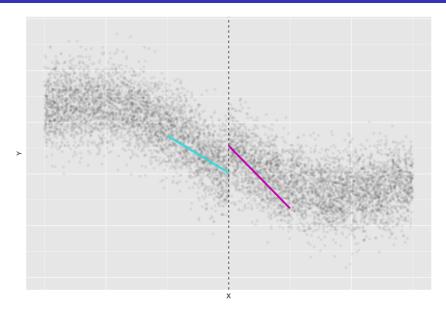
- Regression Discontinuity Donuts:
- Synthetic Controls with Spillovers: Examples and Simulations
- ECDF two-sample package

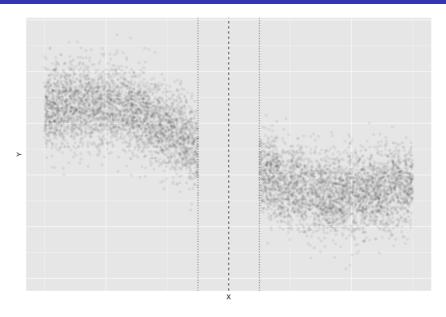
# Regression Discontinuity Donuts

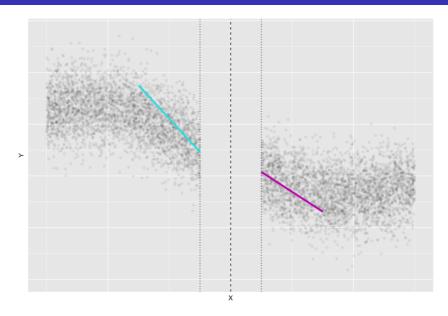
# RD Example

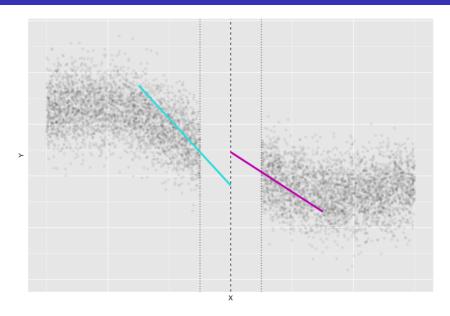


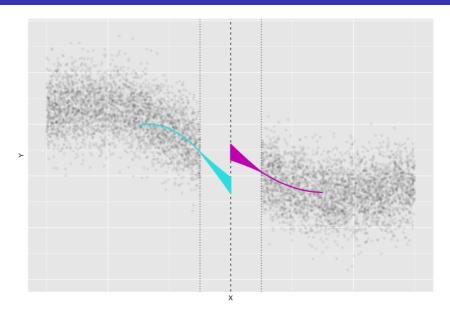
# RD Example











#### Main result

#### Under:

- natural extensions of standard assumptions,
- known or data-determined derivative bounds,
- 3 and straightforward assumptions about selection

we get partial identification for causal effects – and validity while conducting inference for the partially identified set.

#### Outline of Procedure

- **1** Set a confidence level  $\alpha$ , and  $\kappa < \alpha$ .
- ② Estimate the k-1 derivatives of  $\mu_t$  at the edge of the donut.
- **9** Predict  $\mu_t$  at c, using its first k-1 derivatives and a Taylor projection.
- **3** Estimate  $\tau(x_0) = \mu_1(x_0) \mu_0(x_0)$  and build a  $1 \alpha + \kappa$  CI.

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- Estimate  $\tau(x_0) = \mu_1(x_0) \mu_0(x_0)$  and build a  $1 \alpha + \kappa$  CI.
- **5** Find a set  $\mathbb{C}_t$  that contains the  $\mu_t^{(k)}$  with probability  $1 \kappa/2$
- Use the extreme values of  $\mathbb{C}_t$  to find the maximal errors in the Taylor projection above.
- **4** Add those maximal errors for each side to the  $1 \alpha + \kappa$  CI for  $\tau$ .

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#### **Derivative Bounds**

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There is a known k > 0 such that

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For data-driven bounds, we also need to attain the bounds somewhere:

$$\bullet \ \mu_t^{(k)}(x) = I_t \text{ for some } x \in \chi/\mathbb{D}$$

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- $\bullet \ \mu_t^{(k)}(x) = I_t \text{ for some } x \in \chi/\mathbb{D}$

We don't need to know  $l_t$  or  $u_t$ , but we need to be able to estimate them 'well'.

Notice that this condition does not allow other treatment policies with a discontinuity in  $\chi$  which affects Y.

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## Regularity Conditions

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- $\mu_t^{(k+2)}$  is continuous
- **3** The density of X,  $f_x$  is absolutely continuous and bounded away from zero over the region of interest  $\chi$ .
- $\sigma_t^2$ () is positive, bounded away from 0, and has two continuous derivatives.

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- $\sigma_t^2$  () is positive, bounded away from 0, and has two continuous derivatives.
- $\sup_{x \in \chi} \mathbb{E}\left[ |\epsilon_i|^3 exp(|\epsilon_i|) | x_i = x \right] < \infty$ which implies  $\mathbb{E}\left[ |\epsilon_i|^3 exp(|\epsilon_i|) \right] < \infty$ .

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## Local Polynomial Conditions

#### Condition 3: Kernel and Bandwidth for Local Polynomial

- ① The kernel function  $K(\cdot)$  has support (-1,1), outside of which it takes value 0.
- ① The bandwidth  $h=h_n$  is set such that as  $n\to\infty$ ,  $h_n\to0$  and  $nh_n^3\to\infty$ .
- $0 \exists \eta \geq h_n \forall n.$

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We use the kernel  $K_h(x) = K(x/h)/h$ .

These are standard conditions for local polynomial regressions. See Fan, Heckman, and Wand [1995].

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#### **Donut Exclusion**

#### Condition D: Donut Exclusion

- ① There is a known interval  $\mathbb{D} = (d_-, d_+)$  such that all individuals who manipulate are contained to the interval, and would be contained in the counterfactual where they do not manipulate.
- There is only one policy with a threshold relevant to the outcome variable inside the region  $[d_- \epsilon, d_+ + \epsilon]$  for some  $\epsilon > 0$ .

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I define manipulation as the difference between the observed running variable, and the value in the counterfactual where all individuals treatment statuses were fixed in advance.

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# Coverage for $\phi$

Define C such that  $\Phi(C) - \Phi(-C) = 1 - \alpha$ 

$$\mathbb{S}_{1-\alpha} = \left[\hat{\tau}_{\textit{I}} - \textit{C}\hat{\sigma}_{\textit{I}} / \sqrt{\textit{n}}, \ \hat{\tau}_{\textit{u}} + \textit{C}\hat{\sigma}_{\textit{u}} / \sqrt{\textit{n}}\right]$$

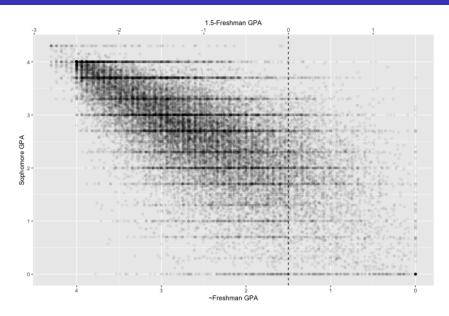
#### Theorem 1

Under conditions 1-4, and the condition that  $nh^{2k+3} \rightarrow 0$ , for all  $\alpha \in (0,1/2)$ ,

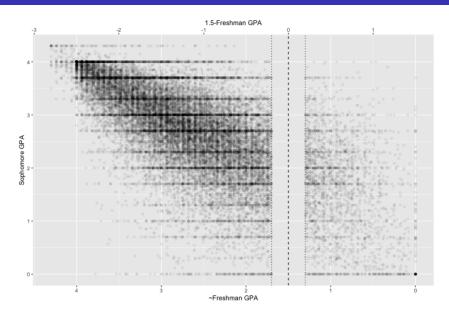
$$\lim_{n\to\infty} P[\phi\subseteq\mathbb{S}_{1-\alpha}]=1-\alpha$$

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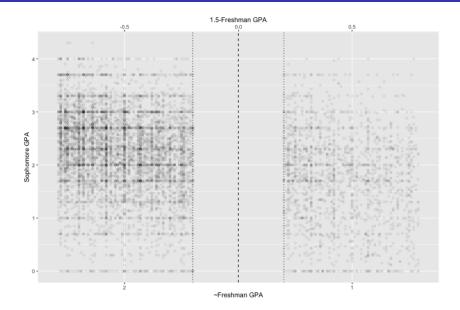
#### Academic Probation - All Data



## Academic Probation - Drop Donut

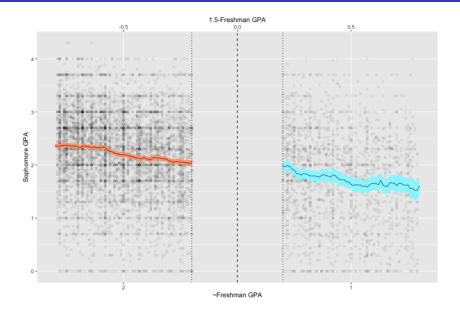


### Academic Probation - Inside Bandwidth



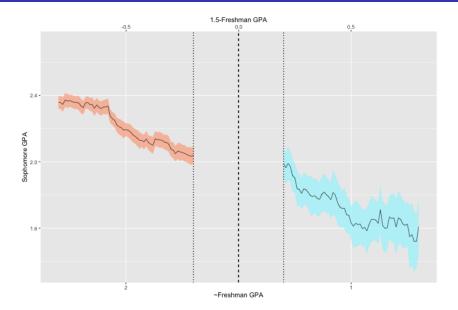
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## Academic Probation - Fit Local Polynomials

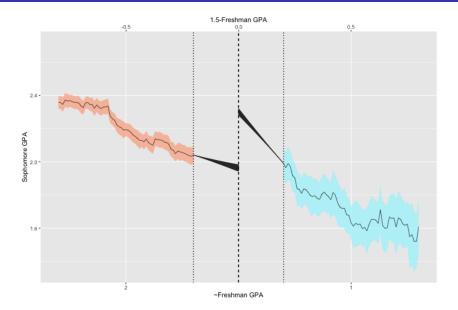


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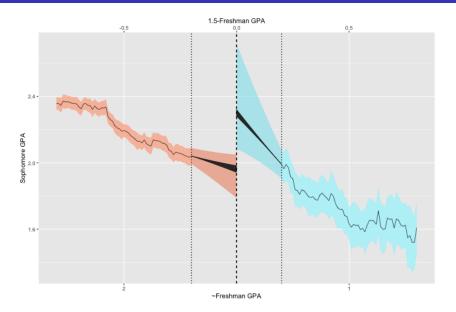
## Academic Probation - Fit Local Polynomials



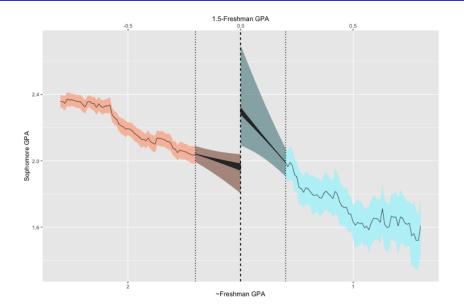
## Academic Probation - Identified Region



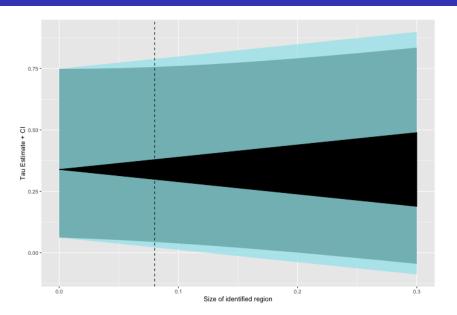
## Academic Probation - CR for Set



## Academic Probation - CR for elements of set

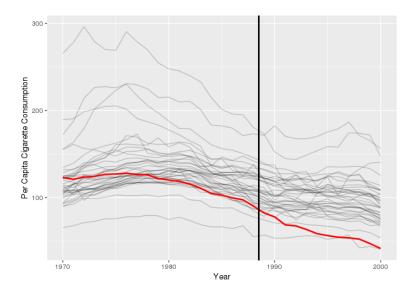


## Tau Set



# Synthetic Controls with Spillovers: Examples and Simulations

## Example: California Prop 9

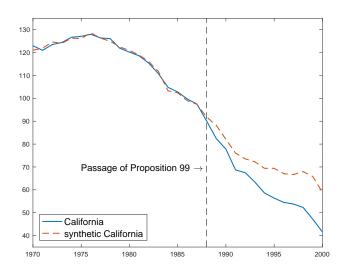


# Synthetic Controls Setting

$$y_{1,T+1}(1) = y_{1,T+1}(0) + \alpha$$

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## Abadie, Diamond, and Hainmueller (2010)



# Synthetic Control Estimator

Synthetic control weights:

$$\begin{pmatrix} \hat{z}_1 \\ \hat{b}_1 \end{pmatrix} = \underset{(z,b')' \in W}{\arg \min} \sum_{t=1}^{T} (y_{1,t} - z - Y_t'b)^2,$$

where 
$$Y_t = (y_{1,t}, y_{2,t}, \dots, y_{N,t})'$$
 and  $W = \mathbb{R} \times \{0\} \times \Delta_{N-1}$ 

i.e.  $\hat{b}_1$  are normalized weights – they sum to one, are non-negative, and we force a weight of 0 on the 'own' observation.

z is included because of the work in Ferman and Pinto (2017) showing an intercept is necessary for unbiased procedures.

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# Spillover effects in synthetic control

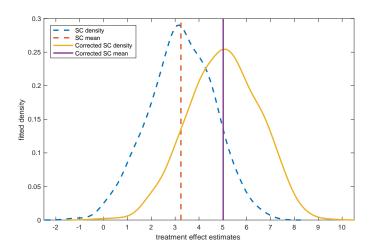
The synthetic control estimator

$$\hat{\alpha} = y_{1,T+1}(1) - \hat{y}_{1,T+1}(0) = y_{1,T+1} - Y'_{T+1}\hat{b}_1 - \hat{z}$$

can be severely biased in the presence of spillover effects:

- As in a diff-in-diff setting, if treatment can affect your control group your ATE may be biased.
- SCM is particularly vulnerable as it may put extra weight on the same units that are 'contaminated'.
- This can be bad luck but more worryingly, similar units may actually be prone to spillovers.
- Regularization properties of simplex are suddenly a potential downside.

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# Linear Spillover effects

Remember:  $\alpha_i = y_{i,T+1}(1) - y_{i,T+1}(0)$  i.e. In the counterfactual, had treatment not occurred anywhere, how different was unit i?

Let

$$\alpha = \begin{bmatrix} \alpha_1 \\ \alpha_2 \\ \vdots \\ \alpha_N \end{bmatrix}$$

Assume linear spillover effects:  $\alpha = A\gamma$ 

- A known
- ullet  $\gamma$  unknown parameters

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# Synthetic control weights for all units

We estimate the weight vectors and intercepts for each unit. For each i,

$$\begin{pmatrix} \hat{z}_i \\ \hat{b}_i \end{pmatrix} = \underset{(z,b) \in W_i}{\operatorname{arg min}} \sum_{t=1}^{T} (y_{i,t} - z - Y_t b')^2,$$

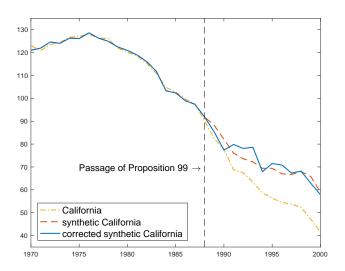
with *i*-th entry of  $\hat{b}_i$  being 0, and  $\hat{b}_i \in \Delta_N$  (non-negative, sum to one).  $Y_t = (y_{1,t}, y_{2,t}, \dots, y_{N,t})'$ 

Define  $z_i = \text{plim } \hat{z}_i$  and  $b_i = \text{plim } \hat{b}_i$ . Let

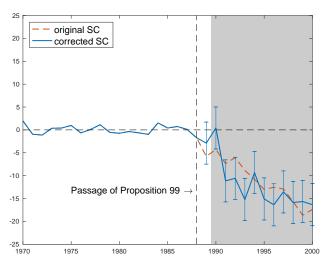
$$Z = \begin{bmatrix} z_1 \\ \vdots \\ z_N \end{bmatrix}, B = \begin{bmatrix} b'_1 \\ \vdots \\ b'_N \end{bmatrix}$$

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# Abadie, Diamond, and Hainmueller (2010)



# Abadie, Diamond, and Hainmueller (2010)



Dark Grey Indicates test for presence of Spillovers rejects Null. Cls are 90%.

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# Two-Sample Tests New test + R package

## Comparing two samples

- Frequently we want to compare two samples, and see if they come from the same distribution.
- Many well known tests are in use for this purpose. T-test, Anova, etc.
- Non-parametric tests, which require few assumptions to work, often are very conservative, or only test for difference in one of a few moments of a distribution.

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#### ECDF test statistics

- ECDF based tests are in this class they are defined by taking some norm on the two empirical cumulative distributions.
- Randomization versions of these tests avoid need to be conservative but are only as powerful as the test statistic allows.

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- I discuss a new ECDF statistic which performs exceptionally well and is motivated by theory.

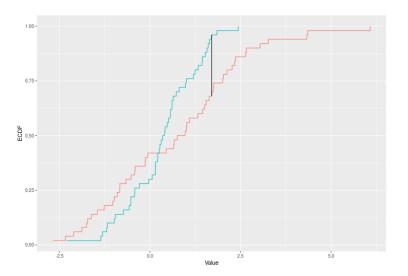
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#### ECDF test statistics

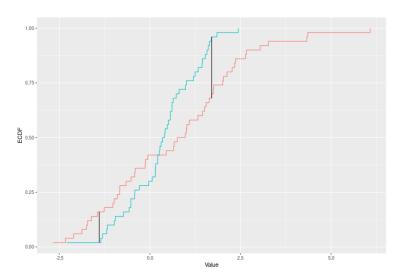
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- Randomization versions of these tests avoid need to be conservative but are only as powerful as the test statistic allows.
- I discuss a new ECDF statistic which performs exceptionally well and is motivated by theory.
- I implement that statistic (and other ECDF tests) in an R package twosamples.

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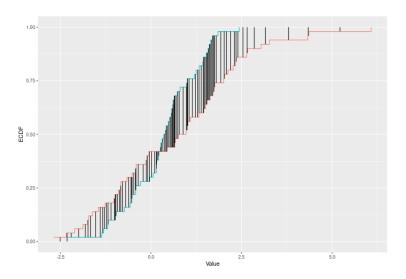
# ECDF Statistics - Kolmogorov-Smirnov



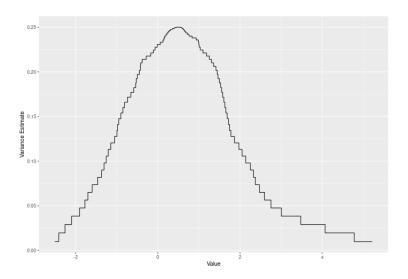
# **ECDF Statistics - Kuiper**



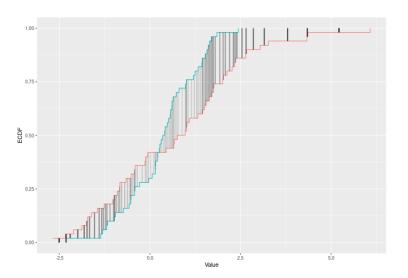
## ECDF Statistics - Cramer-Von Mises



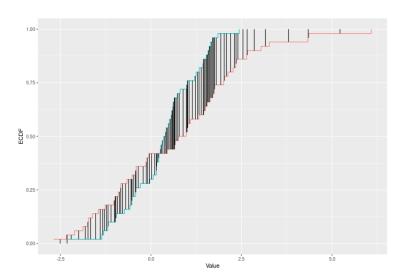
## ECDF Statistics - Variance of ECDF



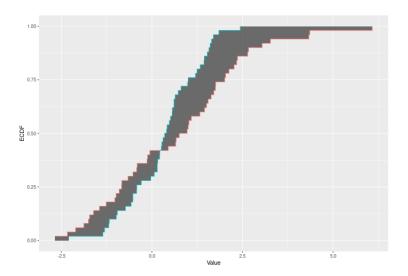
# ECDF Statistics - Anderson-Darling



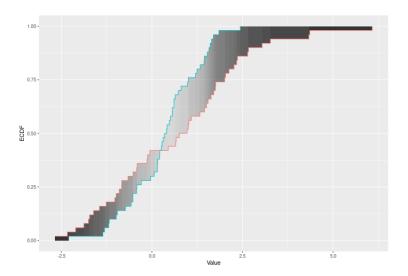
## ECDF Statistics - Cramer-Von Mises



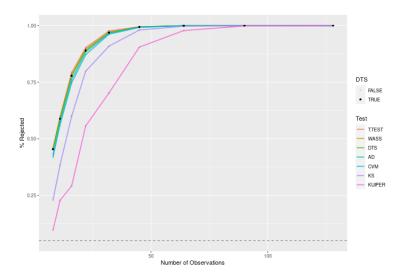
## ECDF Statistics - Wasserstein



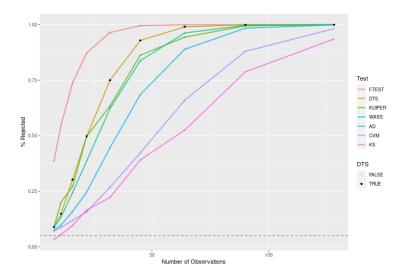
# ECDF Statistics - Proposal



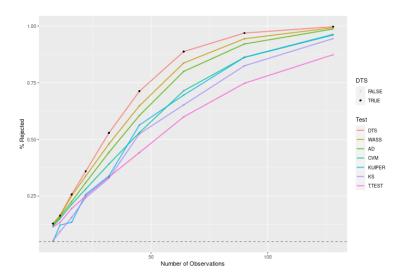
Normal w/ Mean Shift



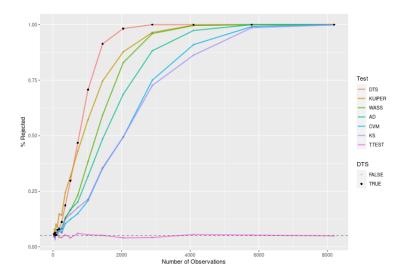
#### Normal w/ Variance Inflation



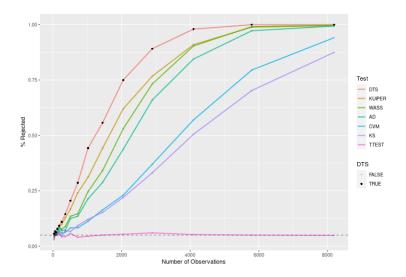
Normal w/ Mean and Variance Shift



#### Mixture of Normals - Mean/Variance Constant - Kurtosis changed



#### Mixture of Normals - Bimodal, Constant mean



#### twosamples

- R package on CRAN
- Publicly available since 2018.
- Randomization tests, with test statistics in C++
- Includes KS, Kuiper, CVM, AD, WASS, and proposed test statistics.
- $\bullet \approx 600 \text{ downloads/month}$

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# Thank you all